

CHAPTER 2: Markov Chains (part 2)

Basic questions

First step analysis

1 Some Markov Chain Models

1.1 An inventory Model

Let X_n denote the number consumable items somebody has at the end of a week. In week n , ξ_n items will be consumed with

$$P(\xi_n = k) = a_k, \quad k = 0, 1, 2, \dots$$

on the weekend, if $X_n > s$, no buying; if $X_n < s$, $S - X_n$ items will be bought. Then

$$X_{n+1} = \begin{cases} X_n - \xi_{n+1}, & \text{if } s < X_n \leq S \\ S - \xi_{n+1}, & \text{if } X_n \leq s. \end{cases}$$

1. It is clear that $\{X_n : n = 0, 1, 2, \dots\}$ is a MC (why?)
2. Let $s = 0$, $S = 2$ with $P(\xi_n = 0) = 0.5$, $P(\xi_n = 1) = 0.4$ and $P(\xi_n = 2) = 0.1$, what is the transition probability matrix ?

$$\mathbf{P} = \begin{matrix} & \begin{matrix} -1 & 0 & 1 & 2 \end{matrix} \\ \begin{matrix} -1 \\ 0 \\ 1 \\ 2 \end{matrix} & \left\| \begin{array}{cccc} 0 & 0.1 & 0.4 & 0.5 \\ 0 & 0.1 & 0.4 & 0.5 \\ 0.1 & 0.4 & 0.5 & 0 \\ 0 & 0.1 & 0.4 & 0.5 \end{array} \right\| \end{matrix}$$

(‘-’ means ‘borrowing’)

3. (unsolved problem) how frequently you need to purchase?

$$\mathbf{P}^\infty = \begin{matrix} & \begin{matrix} -1 & 0 & 1 & 2 \end{matrix} \\ \begin{matrix} -1 \\ 0 \\ 1 \\ 2 \end{matrix} & \left\| \begin{array}{cccc} 0.0444 & 0.2333 & 0.4444 & 0.2778 \\ 0.0444 & 0.2333 & 0.4444 & 0.2778 \\ 0.0444 & 0.2333 & 0.4444 & 0.2778 \\ 0.0444 & 0.2333 & 0.4444 & 0.2778 \end{array} \right\| \end{matrix}$$

we will discuss it later. Note that P^n converges as $n \rightarrow \infty$

1.2 The Ehrenfest Urn Models

suppose there are $2a$ balls. amongst them, k balls are in container (urn) A, and $2a - k$ are in container B. A balls is selected at random (all selections are equally likely) from the totally $2a$ balls and moved to the other container. Let Y_n be the number of balls in container A at the n th stage. Then Y_n is a Markov Chain with

$$S = \{0, 1, 2, \dots, 2a\}.$$

and

$$\begin{aligned} P(Y_{n+1} = i + 1 | Y_n = i) &= \frac{2a - i}{2a}, \text{ if } 0 \leq i \leq 2a \\ P(Y_{n+1} = i - 1 | Y_n = i) &= \frac{i}{2a}, \\ P(Y_{n+1} = j | Y_n = i) &= 0, \text{ if } j \neq i - 1, i + 1. \end{aligned}$$

For $a = 2$, the transition probability matrix is

$$\mathbf{P} = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ 4 \end{matrix} & \left\| \begin{array}{ccccc} 0 & 1 & 0 & 0 & 0 \\ 1/4 & 0 & 3/4 & 0 & 0 \\ 0 & 1/2 & 0 & 1/2 & 0 \\ 0 & 0 & 3/4 & 0 & 1/4 \\ 0 & 0 & 0 & 1 & 0 \end{array} \right\| \end{matrix}$$

We have

$$\mathbf{P}^{2k} = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ 4 \end{matrix} & \left\| \begin{array}{ccccc} 0.125 & 0 & 0.75 & 0 & 0.125 \\ 0 & 0.5 & 0 & 0.5 & 0 \\ 0.125 & 0 & 0.75 & 0 & 0.125 \\ 0 & 0.5 & 0 & 0.5 & 0 \\ 0.125 & 0 & 0.75 & 0 & 0.125 \end{array} \right\| \end{matrix}$$

and

$$\mathbf{P}^{2k+1} = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ 4 \end{matrix} & \left\| \begin{array}{ccccc} 0 & 0.5 & 0 & 0.5 & 0 \\ 0.125 & 0 & 0.75 & 0 & 0.125 \\ 0 & 0.5 & 0 & 0.5 & 0 \\ 0.125 & 0 & 0.75 & 0 & 0.125 \\ 0 & 0.5 & 0 & 0.5 & 0 \end{array} \right\| \end{matrix}$$

Note that P^n does not converge.

1.3 A discrete queueing Markov Chain

Customers arrive for service and take their place in a waiting line¹. Suppose that

$$P(k \text{ customer arrive in a service period})$$

¹A more advanced topic

$$\begin{aligned}
&= P(\xi_n = k) \\
&= a_k.
\end{aligned}$$

where ξ_n has the same distribution as ξ . In each service period, only one customer is served. Let X_n be the customers waiting for service. Then

$$X_{n+1} = \max(X_n - 1, 0) + \xi_n.$$

Based on this, the transition probability matrix is

$$\mathbf{P} = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ \vdots \end{matrix} & \left\| \begin{matrix} a_0 & a_1 & a_2 & a_3 & \dots \\ a_0 & a_1 & a_2 & a_3 & \dots \\ 0 & a_0 & a_1 & a_2 & \dots \\ 0 & 0 & a_0 & a_1 & \dots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \dots \end{matrix} \right\| \end{matrix}$$

Unsolved questions

1. what is \mathbf{P}^n ?
2. Intuitively,
 - (a) If $E\xi > 1$, then the number of customers waiting for service will increase infinitely.
 - (b) If $E\xi < 1$, what is the probability that there will be k customers waiting for service. [if you are the only hairdresser in a barbershop, how many chairs you need to provide?]

2 First Step Analysis

2.1 Motivating Example

In the Gambler's Ruin example (with $\underline{N=4}$ and $X_0 = 3$),

1. what is probability that the gambler eventually goes broke (or win N)?
2. On average, how many games he can play before he goes broke?
3. how many times he can have $0 < k \leq N$ dollars before the game ends?

Recall (for $N = 4$) that

$$\mathbf{P} = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ 4 \end{matrix} & \left\| \begin{matrix} 1 & 0 & 0 & 0 & 0 \\ 2/3 & 0 & 1/3 & 0 & 0 \\ 0 & 2/3 & 0 & 1/3 & 0 \\ 0 & 0 & 2/3 & 0 & 1/3 \\ 0 & 0 & 0 & 0 & 1 \end{matrix} \right\| \end{matrix}$$

States 0 and 4 are absorbing states.

Definition If $P_{ii} = 1$, then state i is an absorbing state.

Two important variables

$$T = \min_n \{n : X_n = 0 \text{ or } X_n = N\} \text{—no. of steps before the end}$$

$$\{X_T = 0\} \text{—}\{\text{the game ends in state 0}\}$$

$$(\text{or } \{X_T = N\} \text{—}\{\text{the game ends in state 0}\}.)$$

By this notation, we go back to the questions

* Question 1 is to calculate $u_3 = P(X_T = 0 | X_0 = 3)$ More generally,

$$u_i = P(X_T = 0 | X_0 = i),$$

where $i = 1, 2, 3$

* Question 2 is to calculate $v_3 = E(T | X_0 = 3)$ More generally,

$$v_i = E(T | X_0 = i), \quad i = 1, 2, 3$$

* Question 3 is to calculate $w_{3k} = E(\sum_{n=0}^{T-1} I(X_n = k) | X_0 = 3)$ More generally,

$$w_{ik} = E\left(\sum_{n=0}^{T-1} I(X_n = k) | X_0 = i\right), i = 1, 2, 3$$

Start from step 1, define

$$T' = \min_n \{n - 1 : X_n = 0 \text{ or } X_n = N\} \text{—no. of steps before the end start with 1}$$

If i is not a absorbing state, then given $X_0 = i$ and $X_1 = j$

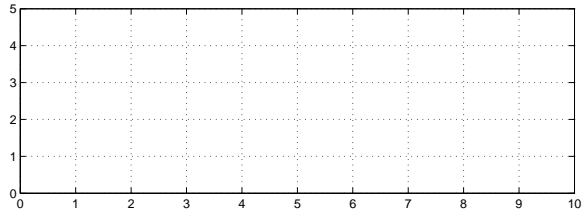
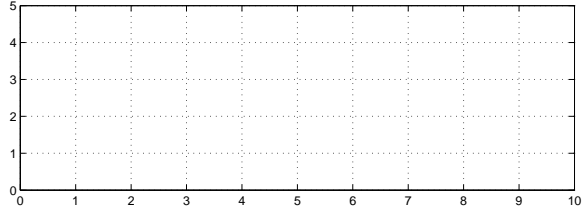
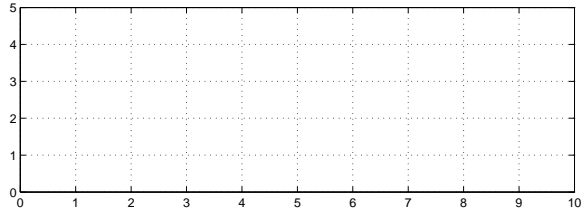
$$T = 1 + T'$$

and given $X_0 = i$, T has the same statistical distribution as T' given $X_1 = i$.

[*Proof: Recall that*

$$P(X_{n_1} = i_{n_1}, \dots, X_{n_k} = i_{n_k} | X_0 = i) = P(X_{n_1+m} = i_{n_1}, \dots, X_{n_k+m} = i_{n_k} | X_m = i)$$

for any $n_1, \dots, n_k > 0$. Thus the above claim follows.]



If state i is not a absorbing state, we have the following useful relations

$$\begin{aligned}
 P(X_T = 0 | X_0 = i, X_1 = j) &= P(X_{T'+1} = 0 | X_0 = i, X_1 = j) \\
 &= P(X_{T'+1} = 0 | X_1 = j) \quad (\text{by MC property}) \\
 &= P(X_T = 0 | X_0 = j) \quad (\text{by the relationship between } T' \text{ and } T) \\
 &= u_j
 \end{aligned}$$

$$\begin{aligned}
 E(T | X_0 = i, X_1 = j) &= 1 + E(T' | X_0 = i, X_1 = j) \\
 &= 1 + E(T' | X_1 = j) \quad (\text{by MC property}) \\
 &= 1 + E(T | X_0 = j) \quad (\text{by the relationship between } T' \text{ and } T) \\
 &= 1 + v_j
 \end{aligned}$$

and

$$\begin{aligned}
 &E\left(\sum_{n=0}^{T-1} I(X_n = k) | X_0 = i, X_1 = j\right) \\
 &= \begin{cases} 1 + E(\sum_{n=1}^{T-1} I(X_n = k) | X_0 = i, X_1 = j) & \text{if } i = k \\ 0 + E(\sum_{n=1}^{T-1} I(X_n = k) | X_0 = i, X_1 = j) & \text{if } i \neq k \end{cases}
 \end{aligned}$$

$$\begin{aligned}
&= \begin{cases} 1 + E(\sum_{n=1}^{T-1} I(X_n = k)|X_1 = j) & \text{if } i = k \\ 0 + E(\sum_{n=1}^{T-1} I(X_n = k)|X_1 = j) & \text{if } i \neq k \end{cases} \\
&= \begin{cases} 1 + E(\sum_{n=1}^{T'+1-1} I(X_n = k)|X_1 = j) & \text{if } i = k \\ 0 + E(\sum_{n=1}^{T'+1-1} I(X_n = k)|X_1 = j) & \text{if } i \neq k \end{cases} \\
&= \begin{cases} 1 + E(\sum_{n=0}^{T-1} I(X_n = k)|X_0 = j) & \text{if } i = k \\ 0 + E(\sum_{n=0}^{T-1} I(X_n = k)|X_0 = j) & \text{if } i \neq k \end{cases} \\
&= \begin{cases} 1 + w_{jk} & \text{if } i = k \\ 0 + w_{jk} & \text{if } i \neq k \end{cases}
\end{aligned}$$

Calculation of u_i in the gambler's example

Let $u_i = P(X_T = 0|X_0 = i), i = 0, 1, 2, 3, 4$. It is easy to see from the example (N=4)

$$u_0 = P(X_T = 0|X_0 = 0) = 1, \quad u_4 = P(X_T = 4|X_0 = 0) = 0$$

By the important relations, we have

$$\begin{aligned}
u_3 &= P(X_T = 0|X_0 = 3) \\
&= \sum_{i=0}^4 P(X_T = 0|X_0 = 3, X_1 = i)P(X_1 = i|X_0 = 3) \\
&= \sum_{i=0}^4 P(X_T = 0|X_0 = i)P(X_1 = i|X_0 = 3) \\
&= \sum_{i=0}^4 u_i P(X_1 = i|X_0 = 3) \\
&= \sum_{i=0}^4 u_i p_{3i}
\end{aligned}$$

More generally,

$$u_j = \sum_{i=0}^4 u_i p_{ji}, \quad j = 1, 2, 3$$

i.e.

$$\begin{aligned}
u_0 &= 1, \\
u_1 &= p_{10}u_0 + p_{11}u_1 + p_{12}u_2 + p_{13}u_3 + p_{14}u_4
\end{aligned}$$

$$\begin{aligned}
u_2 &= p_{20}u_0 + p_{21}u_1 + p_{22}u_2 + p_{23}u_3 + p_{24}u_4 \\
u_3 &= p_{30}u_0 + p_{31}u_1 + p_{32}u_2 + p_{33}u_3 + p_{34}u_4 \\
u_4 &= 0
\end{aligned}$$

or

$$\begin{aligned}
u_0 &= 1 \\
u_1 &= \frac{2}{3}u_0 + 0u_1 + \frac{1}{3}u_2 + 0u_3 + 0u_4 \\
u_2 &= 0u_0 + \frac{2}{3}u_1 + 0u_2 + \frac{1}{3}u_3 + 0u_4 \\
u_3 &= 0u_0 + 0u_1 + \frac{2}{3}u_2 + 0u_3 + \frac{1}{3}u_4 \\
u_4 &= 0
\end{aligned}$$

we have

$$u_1 = \frac{14}{15}, \quad u_2 = \frac{4}{5}, \quad u_3 = \frac{8}{15}$$

Interestingly,

$$\mathbf{P}^n \rightarrow \left\| \begin{array}{ccccc} 1 & 0 & 0 & 0 & 0 \\ 14/15 & 0 & 0 & 0 & 1/15 \\ 4/5 & 0 & 0 & 0 & 1/5 \\ 8/15 & 0 & 0 & 0 & 7/15 \\ 0 & 0 & 0 & 0 & 1 \end{array} \right\|$$

Calculation of $v_i = E(T|X_0 = i)$ in the gambler's example

Note that

$$v_0 = E(T|X_0 = 0) = 0, \quad v_4 = E(T|X_0 = 4) = 0$$

By the important relations

$$\begin{aligned}
v_3 &= E(T|X_0 = 3) \\
&= \sum_{i=0}^4 E(T|X_0 = 3, X_1 = i)E(X_1 = i|X_0 = 3) \\
&= \sum_{i=0}^4 (1 + v_i) \times P(X_1 = i|X_0 = 3) \\
&= \sum_{i=0}^4 p_{3i} + \sum_{i=0}^4 v_i p_{3i}
\end{aligned}$$

Generally

$$v_j = E(T|X_0 = j) = 1 + \sum_{i=0}^4 v_i p_{ji}, \quad j = 1, 2, 3.$$

or

$$\begin{aligned} v_0 &= 0 \\ v_1 &= 1 + \frac{2}{3}v_0 + 0v_1 + \frac{1}{3}v_2 + 0v_3 + 0v_4 \\ v_2 &= 1 + 0v_0 + \frac{2}{3}v_1 + 0v_2 + \frac{1}{3}v_3 + 0v_4 \\ v_3 &= 1 + 0v_0 + 0v_1 + \frac{2}{3}v_2 + 0v_3 + \frac{1}{3}v_4 \\ v_4 &= 0 \end{aligned}$$

we have

$$v_1 = \frac{11}{5}, \quad v_2 = \frac{18}{5}, \quad v_3 = \frac{17}{5}$$

Calculation of $w_{i3} = E(\sum_{n=0}^{T-1} I(X_n = 3)|X_0 = i)$ in the gambler's example

Following the same idea

$$\begin{aligned} w_{33} &= \sum_{j=0}^3 E\left(\sum_{n=0}^{T-1} I(X_n = 3)|X_0 = 3, X_1 = j\right)p_{3j} \\ &= \sum_{j=0}^3 \left\{1 + E\left(\sum_{n=0}^{T-1} I(X_n = 3)|X_0 = j\right)\right\}p_{3j} \\ &= \sum_{j=0}^3 p_{3j} + \sum_{j=0}^3 E\left(\sum_{n=0}^{T-1} I(X_n = 3)|X_0 = j\right)p_{3j} \\ &= 1 + p_{30}w_{03} + p_{31}w_{13} + p_{32}w_{23} + p_{33}w_{33} + p_{34}w_{43} \end{aligned}$$

and similarly for $i \neq 3$

$$w_{i3} = p_{i0}w_{03} + p_{i1}w_{13} + p_{i2}w_{23} + p_{i3}w_{33} + p_{i4}w_{43}$$

special cases

$$w_{03} = 0, w_{43} = 0$$

Combining these, we have

$$\begin{aligned} w_{03} &= 0 \\ w_{13} &= \frac{2}{3}w_{03} + 0w_{13} + \frac{1}{3}w_{23} + 0w_{33} + 0w_{43} \\ w_{23} &= 0w_{03} + \frac{2}{3}w_{13} + 0w_{23} + \frac{1}{3}w_{33} + 0w_{43} \\ w_{33} &= 1 + 0w_{03} + 0w_{13} + \frac{2}{3}w_{23} + 0w_{33} + \frac{1}{3}w_{43} \\ w_{43} &= 0 \end{aligned}$$

we have

$$w_{13} = 0.2, \quad w_{23} = 0.6, \quad w_{33} = 1.4.$$

Another example

Example A Markov chain $X_t : t = 0, 1, \dots$ has the transition probability matrix

$$\mathbf{P} = \begin{array}{c} \begin{array}{ccc} & 0 & 1 & 2 \\ \begin{array}{c} 0 \\ 1 \\ 2 \end{array} & \left\| \begin{array}{ccc} 0.3 & 0.2 & 0.5 \\ 0.5 & 0.1 & 0.4 \\ 0 & 0 & 1 \end{array} \right\| \end{array}$$

and is known to start in state $X_0 = 0$. Eventually, the process will end up in state 2. What is the probability that when the process moves to state 2, it does so from state 1?

[Solution: Let $T = \min\{n \geq 0 : X_n = 2\}$ and

$$z_i = P\{X_{T-1} = 1 | X_0 = i\}, \text{ for } i = 0, 1$$

then we have

$$\begin{aligned} z_0 &= P\{X_{T-1} = 1 | X_0 = 0, X_1 = 0\}p_{00} \\ &\quad + P\{X_{T-1} = 1 | X_0 = 0, X_1 = 1\}p_{01} \\ &\quad + P\{X_{T-1} = 1 | X_0 = 0, X_1 = 2\}p_{02} \\ &= P\{X_{T-1} = 1 | X_0 = 0\}p_{00} + P\{X_{T-1} = 1 | X_0 = 1\}p_{01} + 0 \\ &= z_0p_{00} + z_1p_{01} \end{aligned}$$

i.e.

$$z_0 = 0.3z_0 + 0.2z_1 \tag{2.1}$$

$$\begin{aligned} z_1 &= P\{X_{T-1} = 1 | X_0 = 1, X_1 = 0\}p_{10} \\ &\quad + P\{X_{T-1} = 1 | X_0 = 1, X_1 = 1\}p_{11} \\ &\quad + P\{X_{T-1} = 1 | X_0 = 1, X_1 = 2\}p_{12} \\ &= P\{X_{T-1} = 1 | X_0 = 0\}p_{00} + P\{X_{T-1} = 1 | X_0 = 1\}p_{01} + 0 \\ &= z_0p_{10} + z_1p_{11} + 1 * p_{12} \end{aligned}$$

i.e.

$$z_1 = 0.5z_0 + 0.1z_1 + 0.4 \tag{2.2}$$

It follows from (2.1) and (2.2) that

$$z_0 = 0.1509 \quad z_1 = 0.5283$$